

# Dynamic Risk-Averse Optimization

Andrzej Ruszczyński



University Paris Sud  
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- Utility models
- Mean–risk models
- Measures of risk
- Optimization of measures of risk
- Stochastic dominance constraints
- Introduction to risk–averse dynamic optimization

## Why Probabilistic Models?

- Wealth of results of probability theory
- Connection to real data via statistics
- Universal language (engineering, economics, medicine, ...)
  
- Probability space  $(\Omega, \mathcal{F}, P)$
- Decision space  $\mathcal{X}$
- Random outcome (e.g., cost)  $Z_x(\omega)$ ,  $Z : \mathcal{X} \times \Omega \rightarrow \mathbb{R}$

## Expected Value Model

$$\min_x \mathbb{E}[Z_x] = \int_{\Omega} Z_x(\omega) P(d\omega)$$

It optimizes the outcome **on average** (Law of Large Numbers?)

## What is Risk?

Existence of **unlikely and undesirable** outcomes - high  $Z_x(\omega)$  for some  $\omega$

Expected Utility Models (von Neumann and Morgenstern, 1944)

$$\min_{x \in X} \mathbb{E} [u(Z_x)] \quad \left( = \int_{\Omega} u(Z_x(\omega)) dP(\omega) \right)$$

$u : \mathbb{R} \rightarrow \mathbb{R}$  is a nondecreasing **disutility** function

Rank Dependent Utility (Distortion) Models (Quiggin, 1982; Yaari, 1987)

$$\min_{x \in X} \int_0^1 F_{Z_x}^{-1}(p) dw(p) \quad F_{Z_x}^{-1}(\cdot) - \text{quantile function}$$

$w : [0, 1] \rightarrow \mathbb{R}$  is a nondecreasing **rank dependent utility** function

Existence of utility functions is derived from systems of axioms,  
but in practice they are difficult to elicit

## Axioms of Expected Utility Theory (von Neumann 1944)

$W$  is a **lottery** of  $Z$  and  $V$  with probabilities  $\alpha \in (0, 1)$  and  $(1 - \alpha)$ , if the probability measure  $\mu_W$  induced by  $W$  on  $\mathbb{R}$  is the corresponding convex combination of the probability measures  $\mu_Z$  and  $\mu_V$  of  $Z$  and  $V$ :

$$\mu_W = \alpha\mu_Z + (1 - \alpha)\mu_V.$$

We write the lottery symbolically as

$$W = \alpha Z \oplus (1 - \alpha)V.$$

For law invariant preferences on the space of real random variables, von Neumann introduced the axioms:

**Independence Axiom:** For all  $Z, V, W \in \mathcal{Z}$  one has

$$Z \triangleleft V \implies \alpha Z \oplus (1 - \alpha)W \triangleleft \alpha V \oplus (1 - \alpha)W, \quad \forall \alpha \in (0, 1)$$

**Archimedean Axiom:** If  $Z \triangleleft V \triangleleft W$ , then  $\alpha, \beta \in (0, 1)$  exist such that

$$\alpha Z \oplus (1 - \alpha)W \triangleleft V \triangleleft \beta Z \oplus (1 - \beta)W$$

## Integral Representation

Suppose the total preorder  $\preceq$  on  $\mathcal{Z}$  is law invariant, and satisfies the independence and Archimedean axioms. Then it has an “affine” numerical representation  $U : \mathcal{Z} \rightarrow \mathbb{R}$ :

$$U(\alpha Z \oplus (1 - \alpha)V) = \alpha U(Z) + (1 - \alpha)U(V).$$

If  $\preceq$  is **weakly continuous**, then a continuous and bounded function  $u : \mathbb{R} \rightarrow \mathbb{R}$  exists, such that

$$U(Z) = \mathbb{E}[u(Z)] = \int_{\Omega} u(Z(\omega)) P(d\omega).$$

**New proof by separation theorem** - Dentcheva & R. 2012

In a more general setting, we may consider only r.v. with finite moments, and then **the boundedness condition on  $u(\cdot)$  can be relaxed.**

$$U(Z) = \mathbb{E}[u(Z)] = \int_{\Omega} u(Z(\omega)) P(d\omega)$$

## Monotonicity

The total preorder  $\preceq$  is **monotonic** with respect to the partial order  $\leq$ , if  $Z \leq V \implies Z \preceq V$ .

We focus on  $\mathcal{Z}$  containing integrable random vectors.

## Risk Aversion

A preference relation  $\preceq$  on  $\mathcal{Z}$  is *risk-averse*, if  $\mathbb{E}[Z|\mathcal{G}] \preceq Z$ , for every  $Z \in \mathcal{Z}$  and every  $\sigma$ -subalgebra  $\mathcal{G}$  of  $\mathcal{F}$ .

## Nondecreasing Convex Disutility

Suppose a total preorder  $\preceq$  on  $\mathcal{Z}$  is weakly continuous, monotonic, risk-averse, and satisfies the independence axiom. Then the utility function  $u : \mathbb{R} \rightarrow \mathbb{R}$  is **nondecreasing and convex**.

## Axioms of Dual Utility Theory (Yaari 1987)

Real random variables  $Z_i$ ,  $i = 1, \dots, n$ , are **comonotonic**, if

$$(Z_i(\omega) - Z_i(\omega'))(Z_j(\omega) - Z_j(\omega')) \geq 0$$

for all  $\omega, \omega' \in \Omega$  and all  $i, j = 1, \dots, n$ .

*Dual Independence Axiom:* For all comonotonic random variables  $Z$ ,  $V$ , and  $W$  in  $\mathcal{Z}$  one has

$$Z \triangleleft V \implies \alpha Z + (1 - \alpha)W \triangleleft \alpha V + (1 - \alpha)W, \quad \forall \alpha \in (0, 1)$$

*Dual Archimedean Axiom:* For all comonotonic random variables  $Z$ ,  $V$ , and  $W$  in  $\mathcal{Z}$ , satisfying the relations

$$Z \triangleleft V \triangleleft W,$$

there exist  $\alpha, \beta \in (0, 1)$  such that

$$\alpha Z + (1 - \alpha)W \triangleleft V \triangleleft \beta Z + (1 - \beta)W$$



## Affine Representation

If the total preorder  $\preceq$  on  $\mathcal{Z}$  is law invariant, and satisfies the dual independence and Archimedean axioms, then a numerical representation  $U : \mathcal{Z} \rightarrow \mathbb{R}$  of  $\preceq$  exists, which satisfies for all comonotonic  $Z, V \in \mathcal{Z}$  and all  $\alpha, \beta \in \mathbb{R}_+$  the equation

$$U(\alpha Z + \beta V) = \alpha U(Z) + \beta U(V).$$

## Integral Representation

Suppose  $\mathcal{Z}$  is the set of bounded random variables. If, additionally,  $\preceq$  is continuous in  $\mathcal{L}_1$  and monotonic, then a bounded, nondecreasing, and continuous function  $w : [0, 1] \rightarrow \mathbb{R}_+$  exists, such that

$$U(Z) = \int_0^1 F_Z^{-1}(p) dw(p), \quad Z \in \mathcal{Z}.$$

**Proof by separation** - Dentcheva & R. 2012

$$U(Z) = \int_0^1 F_Z^{-1}(p) dw(p), \quad Z \in \mathcal{Z} \quad (*)$$

## Risk Aversion

A preference relation  $\preceq$  on  $\mathcal{Z}$  is *risk-averse*, if  $\mathbb{E}[Z|\mathcal{G}] \preceq Z$ , for every  $Z \in \mathcal{Z}$  and every  $\sigma$ -subalgebra  $\mathcal{G}$  of  $\mathcal{F}$ .

## Convex Rank-Dependent Utility

Suppose a total preorder  $\preceq$  on  $\mathcal{Z}$  is continuous, monotonic, and satisfies the dual independence axiom. Then it is risk-averse if and only if it has the integral representation (\*) with a **nondecreasing and convex** function  $w : [0, 1] \rightarrow [0, 1]$  such that  $w(0) = 0$  and  $w(1) = 1$ .

## Two Objectives

- Minimize the expected outcome, the **mean**  $\mathbb{E}[Z_x]$
- Minimize a scalar measure of uncertainty of  $Z_x$ , the **risk**  $r[Z_x]$

$$r[Z] = \text{Var}[Z] \quad (\text{Markowitz' model})$$

$$\sigma_p^+[Z] = \left( \mathbb{E}[(Z - \mathbb{E}Z)_+]^p \right)^{1/p} \quad (\text{semideviation})$$

$$\delta_\alpha^+[Z] = \min_{\eta} \mathbb{E} \left[ \max \left( \eta - Z, \frac{\alpha}{1-\alpha} (Z - \eta) \right) \right] \quad (\text{deviation from quantile})$$

## Mean–Risk Optimization

$$\min_{x \in X} \rho[Z_x] = \mathbb{E}[Z_x] + \kappa r[Z_x], \quad 0 \leq \kappa \leq \kappa_{\max}$$

$r[Z_x]$  is **nonlinear w.r.t. probability** and possibly **nonconvex** in  $x$

## Example: Portfolio Optimization

$R_1, R_2, \dots, R_n$  - random return rates of securities

$x_1, x_2, \dots, x_n$  - fractions of the capital invested in the securities

Return rate of the portfolio (negative of)

$$Z_x = -(R_1 x_1 + R_2 x_2 + \dots + R_n x_n)$$

### Risk Optimization with Fixed Mean

$$\begin{aligned} \min_x \quad & r[Z_x] \\ \text{s.t.} \quad & \mathbb{E}[Z_x] = \mu \quad (\text{parameter}) \\ & x \in X_0. \end{aligned}$$

### Combined Mean–Risk Optimization

$$\min_{x \in X_0} \rho[Z_x] = \mathbb{E}[Z_x] + \kappa r[Z_x], \quad 0 \leq \kappa \leq \kappa_{\max}$$

Interesting applications of **parametric optimization**

# Nonlinear Programming Formulations for Discrete Distributions

Suppose  $Z$  has finitely many realizations  $z_1, z_2, \dots, z_S$   
with probabilities  $p_1, p_2, \dots, p_S$

$$\begin{aligned}\rho(Z) &= \mathbb{E}[Z] + \kappa \sigma_m^+[Z] = \mathbb{E}[Z] + \kappa \left( \mathbb{E}[(Z - \mathbb{E}Z)_+]^m \right)^{1/m} \\ &= \sum_{s=1}^S p_s z_s + \kappa \left( \sum_{s=1}^S p_s \left( z_s - \sum_{j=1}^S p_j z_j \right)_+^m \right)^{1/m}\end{aligned}$$

Equivalent Problem (for  $m = 1$  - linear programming)

$$\begin{aligned}\rho(Z) &= \min_{v, \mu} \quad \mu + \kappa \left( \sum_{s=1}^S p_s v_s^m \right)^{1/m} \\ \text{s.t.} \quad &\mu = \sum_{s=1}^S p_s z_s \\ &v_s \geq z_s - \mu, \quad s = 1, \dots, S \\ &v_s \geq 0, \quad s = 1, \dots, S\end{aligned}$$

Suppose the vector of return rates has  $S$  realizations with probabilities  $p_1, p_2, \dots, p_S$

$R_{js}$  - return rate of asset  $j = 1, \dots, n$  in scenario  $s = 1, \dots, S$

Equivalent Problem (for  $m = 1$  - linear programming)

$$\min_{x, z, v, \mu} \quad \mu + \kappa \left( \sum_{s=1}^S v_s^m \right)^{1/m}$$

$$\text{s.t.} \quad \mu = \sum_{s=1}^S p_s z_s$$

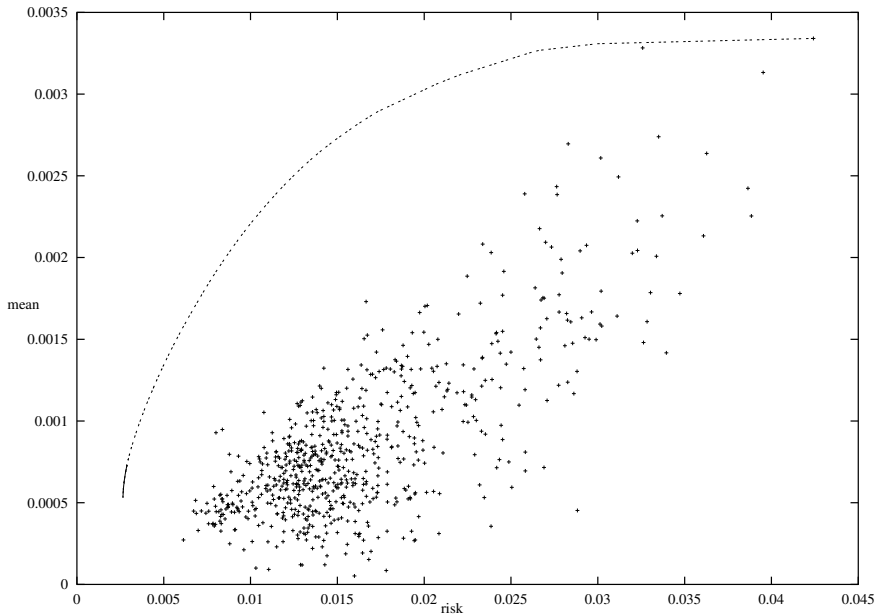
$$z_s = - \sum_{j=1}^n R_{sj} x_j, \quad s = 1, \dots, S$$

$$v_s \geq z_s - \mu, \quad s = 1, \dots, S$$

$$v_s \geq 0, \quad s = 1, \dots, S$$

$$x \in X_0$$

# Mean–Semideviation Model (719 stocks); (with R. Vanderbei)



## Key Requirement: Monotonicity

$$\rho(Z) = \mathbb{E}[Z] + \kappa r[Z]$$

Consistency with Stochastic Dominance (Ogryczak–R., 1997)

$$\mathbb{E}[u(Z)] \leq \mathbb{E}[u(W)], \forall \text{ nondecreasing and convex } u(\cdot) \Rightarrow \rho[Z] \leq \rho[W]$$

Consistency with Pointwise Order (Artzner *et al.*, 1999)

$$Z \leq W \text{ a.s.} \Rightarrow \rho[Z] \leq \rho[W]$$

Mean–semideviation and mean–deviation from quantile models are consistent for  $0 \leq \kappa \leq 1$ , but **not mean–variance**.

Unique optimal solutions of consistent optimization models

$$\min_{x \in X} \rho(Z_x)$$

cannot be strictly dominated (in the corresponding sense)



Space of uncertain outcomes  $\mathcal{Z} = \mathcal{L}_p(\Omega, \mathcal{F}, P)$ ,  $p \in [1, \infty]$

A functional  $\rho : \mathcal{Z} \rightarrow \overline{\mathbb{R}}$  is a **coherent risk measure** if it satisfies the following axioms

- **Convexity:**  $\rho(\lambda Z + (1 - \lambda)W) \leq \lambda\rho(Z) + (1 - \lambda)\rho(W)$   
 $\forall \lambda \in (0, 1), Z, W \in \mathcal{Z}$
- **Monotonicity:** If  $Z \leq W$  then  $\rho(Z) \leq \rho(W)$ ,  $\forall Z, W \in \mathcal{Z}$
- **Translation Equivariance:**  $\rho(Z + a) = \rho(Z) + a$ ,  $\forall Z \in \mathcal{Z}, a \in \mathbb{R}$
- **Positive Homogeneity:**  $\rho(\tau Z) = \tau\rho(Z)$ ,  $\forall Z \in \mathcal{Z}, \tau \geq 0$

Kijima-Ohnishi (1993) – no monotonicity

Artzner-Delbaen-Eber-Heath (1999–) - space  $\mathcal{L}_\infty$

R.-Shapiro (2005) – spaces  $\mathcal{L}_p, \dots$

**Good news:**  $\mathbb{E}[Z]$  is coherent

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The Value at Risk at level  $\alpha \in (0, 1)$  of a random cost  $Z \in \mathcal{Z}$ :

$$\text{VaR}_\alpha^+(Z) \triangleq \inf \{ \eta : F_Z(\eta) \geq 1 - \alpha \} = F_Z^{-1}(1 - \alpha)$$

**Monotonicity:**  $Z \leq V \implies \text{VaR}_\alpha^+(Z) \leq \text{VaR}_\alpha^+(V)$

**Translation:**  $\text{VaR}_\alpha^+(Z + c) = \text{VaR}_\alpha^+(Z) + c$ , for all  $c \in \mathbb{R}$

**Positive Homogeneity:**  $\text{VaR}_\alpha^+(\gamma Z) = \gamma \text{VaR}_\alpha^+(Z)$ , for all  $\gamma \geq 0$

However, **it is not convex**

**Counterexample: Two independent variables**

$$Z = \begin{cases} 0 & \text{with probability } 1 - p \\ 1 & \text{with probability } p \end{cases} \quad V = \begin{cases} 0 & \text{with probability } 1 - p \\ 1 & \text{with probability } p \end{cases}$$

For  $p < \alpha < 1$  we have  $\text{VaR}_\alpha^+(Z) = \text{VaR}_\alpha^+(V) = 0$

If  $p < \alpha < 1 - (1 - p)^2$ , we have **non-convexity**

$$\text{VaR}_\alpha^+(\lambda Z + (1 - \lambda)V) > 0 = \lambda \text{VaR}_\alpha^+(Z) + (1 - \lambda) \text{VaR}_\alpha^+(V)$$

$$\text{AV@R}_\alpha^+(Z) \triangleq \frac{1}{\alpha} \int_0^\alpha \text{V@R}_\beta^+(Z) \, d\beta$$

If the  $(1 - \alpha)$ -quantile of  $Z$  is unique

$$\text{AV@R}_\alpha^+(Z) = \frac{1}{\alpha} \int_{\text{V@R}_\alpha^+(Z)}^\infty z \, dF_Z(z) = \mathbb{E}[Z | Z \geq \text{V@R}_\alpha^+(Z)]$$

### Extremal representation

$$\text{AV@R}_\alpha^+(Z) = \inf_{\eta \in \mathbb{R}} \left\{ \eta + \frac{1}{\alpha} \mathbb{E}[(Z - \eta)_+] \right\}$$

The minimizer  $\eta = \text{V@R}_\alpha(Z)$

Connection to weighted deviation from  $\alpha$ -quantile:

$$\delta_\alpha^+(Z) = \text{AV@R}_\alpha^+(Z) - \mathbb{E}[Z], \quad \alpha \in [0, 1].$$



Suppose  $Z$  has finitely many realizations  $z_1, z_2, \dots, z_S$  with probabilities  $p_1, p_2, \dots, p_S$

$$\begin{aligned} \min_{v, \eta} \quad & \eta + \frac{1}{\alpha} \sum_{s=1}^S p_s v_s \\ \text{s.t.} \quad & v_s \geq z_s - \eta, \quad s = 1, \dots, S \\ & v_s \geq 0, \quad s = 1, \dots, S \end{aligned}$$

For portfolios we have to add the constraints

$$\begin{aligned} z_s &= - \sum_{j=1}^n R_{sj} x_j, \quad s = 1, \dots, S \\ x &\in X_0 \end{aligned}$$

and include  $z$  and  $x$  into the decision variables

# Conjugate Duality of Risk Measures

**Pairing** of a linear topological space  $\mathcal{Z}$  with a linear topological space  $\mathcal{Y}$  of regular signed measures on  $\Omega$  with the bilinear form

$$\langle \mu, Z \rangle = \mathbb{E}_\mu[Z] = \int_\Omega Z(\omega) \mu(d\omega)$$

We assume standard conditions on pairing and the polarity:  $(\mathcal{Z}_+)^{\circ} = \mathcal{Y}_-$

## Dual Representation Theorem

If  $\rho : \mathcal{Z} \rightarrow \overline{\mathbb{R}}$  is a lower semicontinuous\* coherent risk measure, then

$$\rho(Z) = \max_{\mu \in \mathcal{A}} \int_\Omega Z(\omega) \mu(d\omega), \quad \forall Z \in \mathcal{Z}$$

with a convex closed  $\mathcal{A} \subset \mathcal{P}$  (set of probability measures in  $\mathcal{Y}$ ).

Delbaen (2001), Föllmer–Schied (2002), R.–Shapiro (2005),

Rockafellar–Uryasev–Zabarankin (2006), ...

\* Lower semicontinuity is automatic if  $\rho$  is finite and  $\mathcal{Z}$  is a Banach lattice

# Universality of AV@R

$Z \sim V$  means that  $Z$  and  $V$  have the same distribution,  $\mu_Z = \mu_V$ .

$\rho : \mathcal{Z} \rightarrow \mathbb{R}$  is **law invariant** if  $Z \sim V \implies \rho(Z) = \rho(V)$

## Kusuoka Theorem

If  $(\Omega, \mathcal{F}, P)$  is **atomless** and  $\rho : \mathcal{L}_1(\Omega, \mathcal{F}, P) \rightarrow \mathbb{R}$  is **law invariant** and **coherent**, then

$$\rho(Z) = \sup_{m \in \mathcal{M}} \int_0^1 \text{AV@R}_\alpha^+(Z) m(d\alpha)$$

where  $\mathcal{M}$  is a convex set of probability measures on  $(0, 1]$ .

## Spectral measure

$$\rho(V) = \int_0^1 \text{AV@R}_\alpha^+(Z) m(d\alpha)$$

Spectral measures have dual utility form:

$$\rho(Z) = \int_0^1 F_Z^{-1}(\beta) d\omega(\beta)$$

# Optimization of Risk Measures

“Minimize” over  $x \in X$  a random outcome  $Z_x(\omega) = f(x, \omega)$ ,  $\omega \in \Omega$

## Composite Optimization Problem

$$\min_{x \in X} \rho(Z_x) \quad (\text{P})$$

## Theorem

Let  $x \mapsto Z_x(\omega)$  be convex and  $\rho(\cdot)$  be coherent. Suppose  $\hat{x} \in X$  is an optimal solution of (P) and  $\rho(\cdot)$  is continuous at  $Z_{\hat{x}}$ . Then there exists a probability measure  $\hat{\mu} \in \partial\rho(Z_{\hat{x}}) \subseteq \mathcal{A}$  such that  $\hat{x}$  solves

$$\min_{x \in X} \mathbb{E}_{\hat{\mu}}[Z_x] = \min_{x \in X} \max_{\mu \in \mathcal{A}} \mathbb{E}_{\mu}[Z_x]$$

We also have the **duality relation**:

$$\min_{x \in X} \rho(Z_x) = \max_{\mu \in \mathcal{A}} \inf_{x \in X} \mathbb{E}_{\mu}[Z_x]$$

## Duality in Portfolio Optimization - Game Model

Suppose the vector of return rates of assets has  $S$  realizations

$R_{js}$  - return rate of asset  $j = 1, \dots, n$  in scenario  $s = 1, \dots, S$

Portfolio return (negative) in scenario  $s$

$$Z_s(x) = - \sum_{j=1}^n R_{js} x_j$$

### Risk-Averse Portfolio Problem

$$\min_{x \in X} \rho(Z(x))$$

By homogeneity, we may assume that  $\sum_{j=1}^n x_j = 1$

### Equivalent Matrix Game

$$\max_{x \in X} \min_{\mu \in \mathcal{A}} \sum_{j=1}^n \sum_{s=1}^S x_j R_{js} \mu_s$$

$x$  - mixed strategy of the investor

$\mu$  - mixed strategy of the opponent (market)

$Z_x$  - random outcome (e.g., cost)

$Y$  - **benchmark** random outcome, e.g.  $Y(\omega) = Z_{\bar{x}}(\omega)$  for some  $\bar{x} \in X$

## New Model

$$\begin{array}{ll} \min \mathbb{E}[Z_x] & \text{(or some other objective)} \\ \text{subject to } Z_x \leq_{\mathcal{U}} Y & \text{(stochastic ordering constraint)} \\ x \in X \end{array}$$

$Z_x$  is preferred over  $Y$  by all decision makers having disutility functions in the generator  $\mathcal{U}$ :

$$\mathbb{E}[u(Z_x)] \leq \mathbb{E}[u(Y)] \quad \forall u \in \mathcal{U}$$

All nondecreasing  $u(\cdot)$  - **first order stochastic dominance**  $\leq_{\text{st}}$

All nondecreasing convex  $u(\cdot)$  - **increasing convex order**  $\leq_{\text{icx}}$

$$\begin{aligned} & \min \mathbb{E}[Z_x] \\ & \text{subject to } Z_x \leq_{\text{icx}} Y \\ & x \in X \end{aligned}$$

$X$  - convex set in  $\mathcal{X}$  (separable locally convex Hausdorff vector space)

$x \mapsto Z_x$  is a continuous operator from  $\mathcal{X}$  to  $\mathcal{L}_1(\Omega, \mathcal{F}, P)$

$x \mapsto Z_x(\omega)$  is convex for  $P$ -almost all  $\omega \in \Omega$

Primal:  $\mathbb{E}[u(Z_x)] \leq \mathbb{E}[u(Y)]$  for all convex nondecreasing  $u : \mathbb{R} \rightarrow \mathbb{R}$

Inverse:  $\int_0^1 F_{Z_x}^{-1}(p) dw(p) \leq \int_0^1 F_Y^{-1}(p) dw(p)$  for all convex nondecreasing  $w : [0, 1] \rightarrow \mathbb{R}$

## Main Results

- Utility functions  $u : \mathbb{R} \rightarrow \mathbb{R}$  and rank dependent utility functions  $w : [0, 1] \rightarrow \mathbb{R}$  play the roles of Lagrange multipliers
- Expected utility models and rank dependent utility models are Lagrangian relaxations of the problem

## Lagrangian in Direct Form

$$L(x, u) = \mathbb{E}[Z_x + u(Z_x) - u(Y)]$$

$u(\cdot)$  - convex function on  $\mathbb{R}$

## Theorem

Assume Uniform Dominance Condition (a form of Slater constraint qualification). If  $\hat{x}$  is an optimal solution of the problem then there exists a function  $\hat{u} \in \mathcal{U}$  such that

$$L(\hat{x}, \hat{u}) = \min_{x \in X} L(x, \hat{u}) \quad (1)$$

$$\mathbb{E}[\hat{u}(Z_{\hat{x}})] = \mathbb{E}[\hat{u}(Y)] \quad (2)$$

Conversely, if for some function  $\hat{u} \in \mathcal{U}$  an optimal solution  $\hat{x}$  of (1) satisfies the dominance constraint and (2), then  $\hat{x}$  is optimal



## Lagrangian in Inverse Form

$$\Phi(x, w) = \int_0^1 F_{Z_x}^{-1}(p) d(p + w(p)) - \int_0^1 F_Y^{-1}(p) dw(p)$$

$w(\cdot)$  - convex function on  $[0, 1]$

## Theorem

Assume Uniform Dominance Condition (a form of Slater constraint qualification). If  $\hat{x}$  is an optimal solution of the problem, then there exists a function  $\hat{w} \in \mathcal{W}$  such that

$$\Phi(\hat{x}, \hat{w}) = \min_{x \in X} \Phi(x, \hat{w}) \tag{3}$$

$$\int_0^1 F_{Z_{\hat{x}}}^{-1}(p) d\hat{w}(p) = \int_0^1 F_Y^{-1}(p) d\hat{w}(p) \tag{4}$$

If for some  $\hat{w} \in \mathcal{W}$  an optimal solution  $\hat{x}$  of (3) satisfies the inverse dominance constraint and (4), then  $\hat{x}$  is optimal

# How to Measure Risk of Sequences?

Probability space  $(\Omega, \mathcal{F}, P)$  with filtration  $\mathcal{F}_1 \subset \dots \subset \mathcal{F}_T \subset \mathcal{F}$

Adapted sequence of random variables (costs)  $Z_1, Z_2, \dots, Z_T$

Spaces:  $\mathcal{Z}_t = \mathcal{L}_{\bar{s}}(\Omega, \mathcal{F}_t, P)$ ,  $\bar{s} \in [1, \infty)$ , and  $\mathcal{Z}_{t,T} = \mathcal{Z}_t \times \dots \times \mathcal{Z}_T$

## Conditional Risk Measure

A mapping  $\rho_{t,T} : \mathcal{Z}_{t,T} \rightarrow \mathcal{Z}_t$  satisfying the **monotonicity condition**:

$$\rho_{t,T}(Z) \leq \rho_{t,T}(W) \text{ for all } Z, W \in \mathcal{Z}_{t,T} \text{ such that } Z \leq W$$

## Dynamic Risk Measure

A sequence of conditional risk measures  $\rho_{t,T} : \mathcal{Z}_{t,T} \rightarrow \mathcal{Z}_t$ ,  $t = 1, \dots, T$

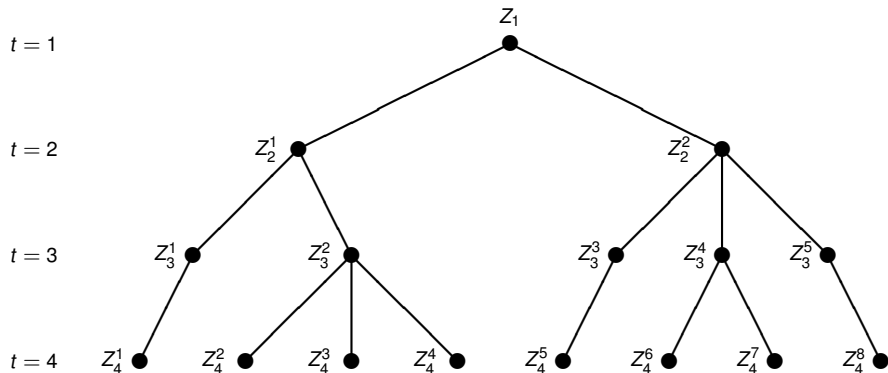
$$\rho_{1,T}(Z_1, Z_2, Z_3, \dots, Z_T) \in \mathcal{Z}_1 = \mathbb{R}$$

$$\rho_{2,T}(Z_2, Z_3, \dots, Z_T) \in \mathcal{Z}_2$$

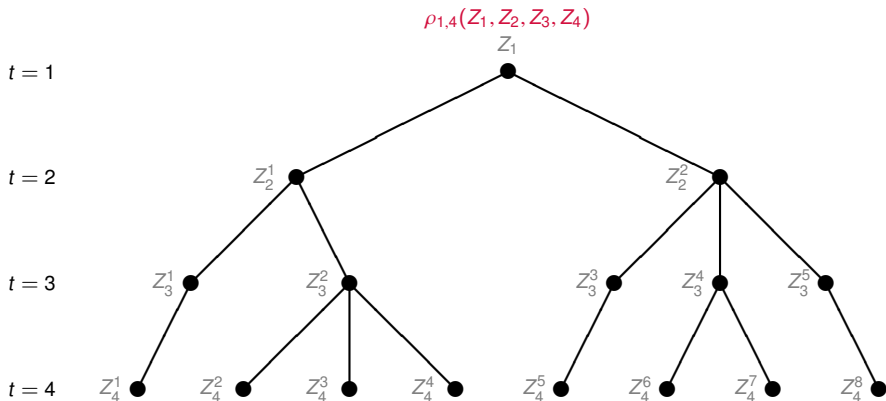
$$\rho_{3,T}(Z_3, \dots, Z_T) \in \mathcal{Z}_3$$

⋮

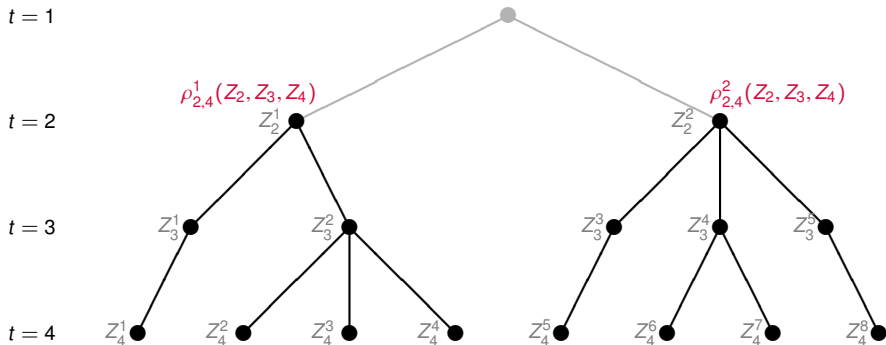
# Evaluating Risk on a Scenario Tree



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- Dynamic measures of risk
- Time consistency and local property
- Interchangeability
- Risk optimization on a tree
- Application to Markov models
- Stochastic conditional time-consistency
- Markov risk measures
- Dynamic programming
- Solution methods
- Examples

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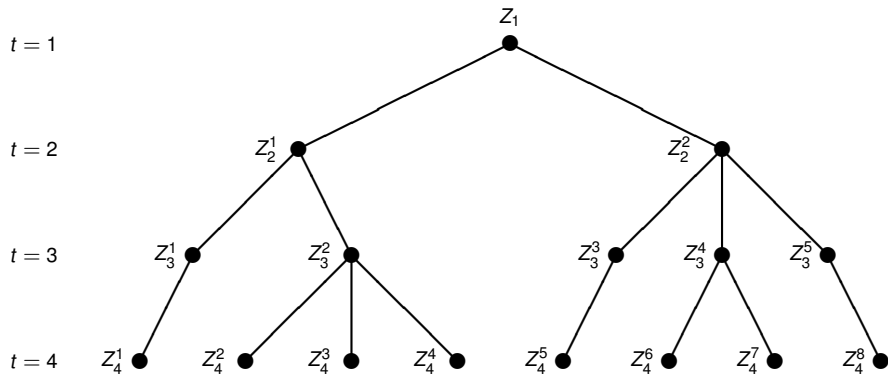
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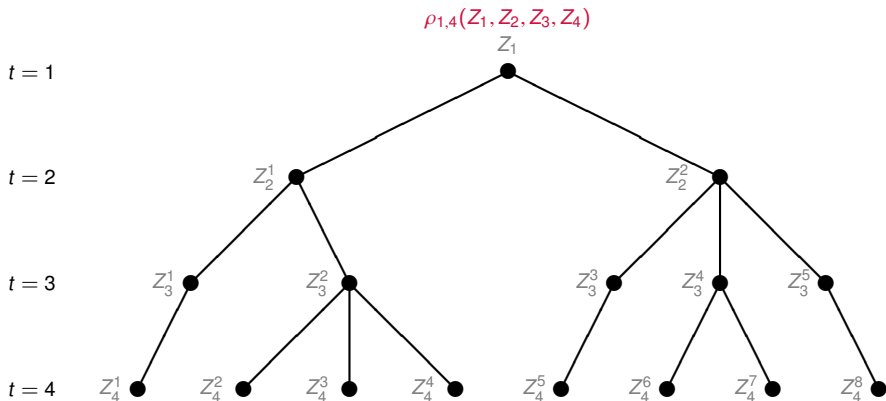
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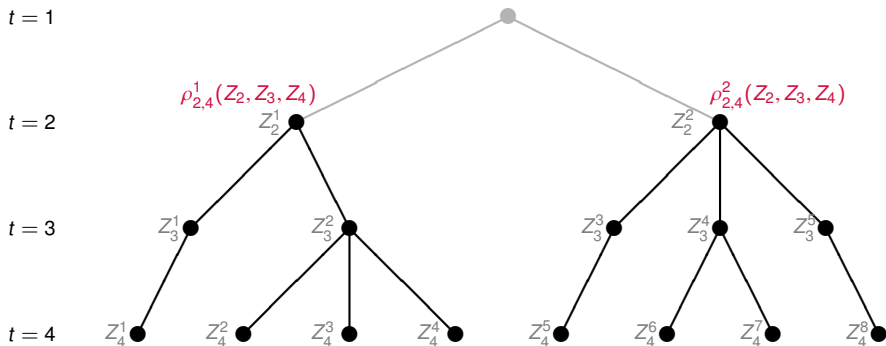
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# Time Consistency of Dynamic Risk Measures

A dynamic risk measure  $\{\rho_{t,T}\}_{t=1}^T$  is **time-consistent** if for all  $\tau < \theta$

$$Z_k = W_k, \quad k = \tau, \dots, \theta - 1 \quad \text{and} \quad \rho_{\theta,T}(Z_\theta, \dots, Z_T) \leq \rho_{\theta,T}(W_\theta, \dots, W_T)$$

imply that  $\rho_{\tau,T}(Z_\tau, \dots, Z_T) \leq \rho_{\tau,T}(W_\tau, \dots, W_T)$

Define  $\rho_t(Z_{t+1}) = \rho_{t,T}(0, Z_{t+1}, 0, \dots, 0)$

## Nested Decomposition Theorem

Suppose a dynamic risk measure  $\{\rho_{t,T}\}_{t=1}^T$  is time-consistent and

$$\rho_{t,T}(Z_t, Z_{t+1}, \dots, Z_T) = Z_t + \rho_{t,T}(0, Z_{t+1}, \dots, Z_T)$$

Then for all  $t$  we have the representation

$$\rho_{t,T}(Z_t, \dots, Z_T) = Z_t + \rho_t \left( Z_{t+1} + \rho_{t+1} \left( Z_{t+2} + \dots + \rho_{T-1}(Z_T) \right) \dots \right)$$

# Coherent One-Step Conditional Risk Measures

Stronger assumptions about one-step measures  $\rho_t : \mathcal{Z}_{t+1} \rightarrow \mathcal{Z}_t$ :

- **Convexity:**  $\rho_t(\lambda Z + (1 - \lambda)W) \leq \lambda \rho_t(Z) + (1 - \lambda)\rho_t(W)$   
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## Example: Conditional Mean–Semideviation

$$\rho_t(Z_{t+1}) = \mathbb{E}[Z_{t+1} | \mathcal{F}_t] + \kappa \mathbb{E} \left[ \left( Z_{t+1} - \mathbb{E}[Z_{t+1} | \mathcal{F}_t] \right)_+^s \middle| \mathcal{F}_t \right]^{\frac{1}{s}}$$

Here  $s \in [1, \bar{s}]$  and  $\kappa \in [0, 1]$  may be  $\mathcal{F}_t$ -measurable

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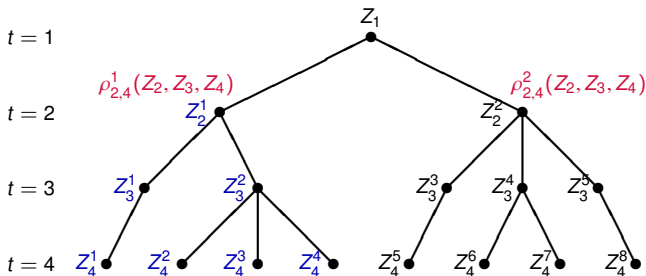
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# Local Property

A conditional risk measure  $\rho_{t,T} : \mathcal{Z}_{t,T} \rightarrow \mathcal{Z}_t$  has the **local property**, if for every event  $A \in \mathcal{F}_t$  we have the equation

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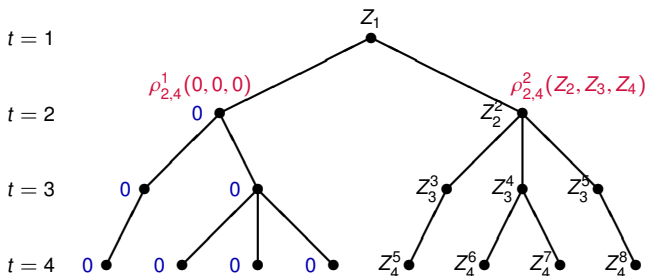


Automatic for **coherent** conditional risk measures

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# Multistage Risk-Averse Optimization Problems

**Probability Space:**  $(\Omega, \mathcal{F}, P)$  with filtration  $\mathcal{F}_1 \subset \dots \subset \mathcal{F}_T \subset \mathcal{F}$

**Decision Variables:**  $x_t(\omega)$ ,  $\omega \in \Omega$ ,  $t = 1, \dots, T$

**Nonanticipativity:** Each  $x_t$  is  $\mathcal{F}_t$ -measurable

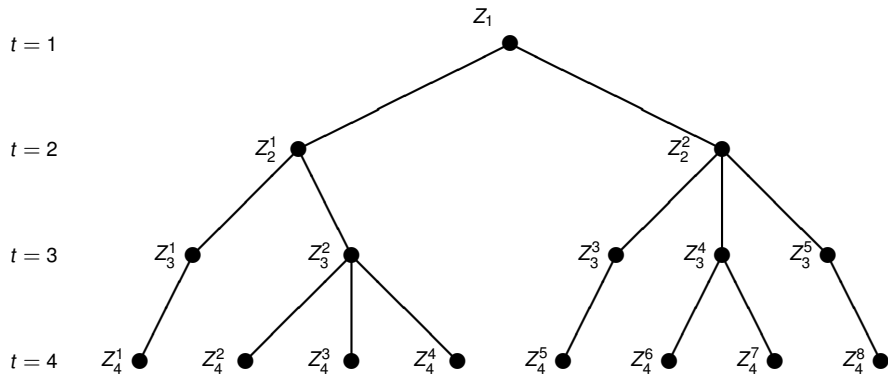
**Cost per Stage:**  $Z_t(x_t)$  with realizations  $Z_t(x_t(\omega), \omega)$ ,  $\omega \in \Omega$

**Objective Function:** Time-consistent dynamic measure of risk

## Interchangeability for Time-Consistent Measures

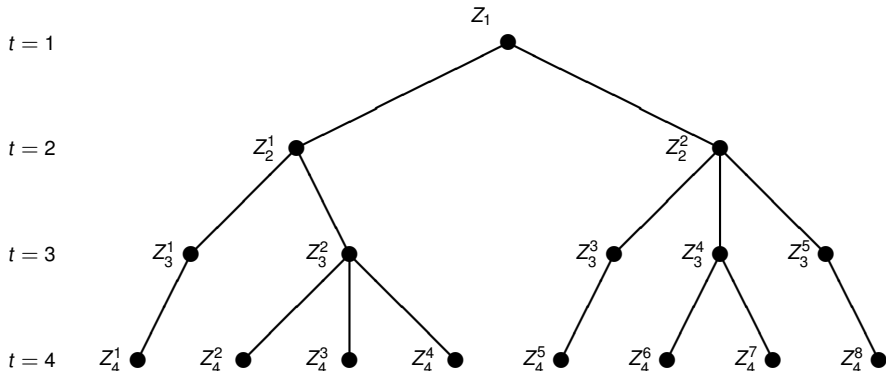
$$\begin{aligned} & \min_{x_1, x_2(\cdot), \dots, x_T(\cdot)} \left\{ Z_1(x_1) + \rho_1 \left( Z_2(x_2) + \rho_2 \left( Z_3(x_3) + \dots \right. \right. \right. \\ & \qquad \qquad \qquad \left. \left. \left. \dots + \rho_{T-2} \left( Z_{T-1}(x_{T-1}) + \rho_{T-1} (Z_T(x_T)) \right) \dots \right) \right) \right\} \\ &= \min_{x_1} \left\{ Z_1(x_1) + \rho_1 \left[ \min_{x_2} \left( Z_2(x_2) + \rho_2 \left[ \min_{x_3} \left( Z_3(x_3) + \dots \right. \right. \right. \right. \right. \right. \\ & \qquad \qquad \qquad \left. \left. \left. \left. \dots + \rho_{T-2} \left[ \min_{x_{T-1}} \left( Z_{T-1}(x_{T-1}) + \rho_{T-1} \left( \min_{x_T} Z_T(x_T) \right) \right) \right] \dots \right) \right] \right] \right] \right\} \end{aligned}$$

# Interchangeability on a Scenario Tree

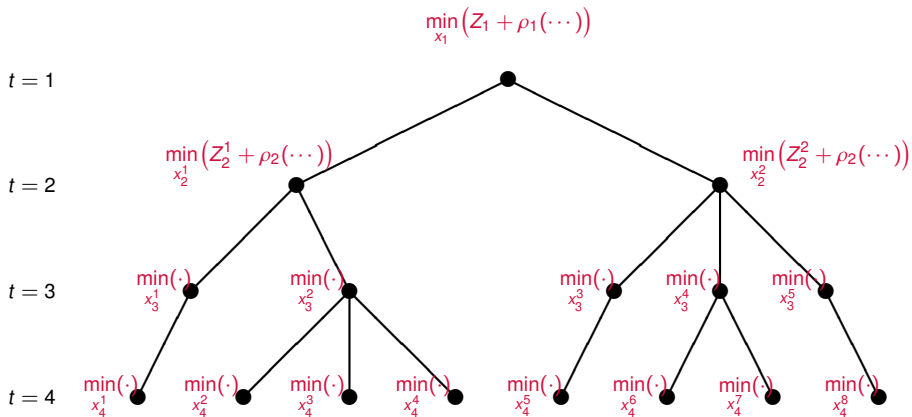


# Interchangeability on a Scenario Tree

$$\min_{x_1, x_2(\cdot), x_3(\cdot), x_4(\cdot)} \rho_{1,4}(Z_1, Z_2, Z_3, Z_4)$$



# Interchangeability on a Scenario Tree





# Linear Risk-Averse Multistage Optimization

$(\Omega, \mathcal{F}, P)$  - probability space with filtration  $\{\emptyset, \Omega\} = \mathcal{F}_1 \subset \mathcal{F}_2 \subset \dots \subset \mathcal{F}_T = \mathcal{F}$ .

A random  $x = (x_1, \dots, x_T)$  is a **policy**.

If each  $x_t$  is  $\mathcal{F}_t$ -measurable, policy  $x$  is **implementable** (belongs to  $I$ ).

A policy  $x$  is **feasible** (belongs to  $F$ ), if it satisfies the conditions:

$$\begin{array}{rcl} A_1 x_1 & & = b_1 \\ B_2 x_1 + A_2 x_2 & & = b_2 \\ \dots & & \dots \\ & B_T x_{T-1} + A_T x_T & = b_T \\ x_1 \in X_1, \quad x_2 \in X_2, \quad \dots \quad x_T \in X_T & & \end{array}$$

Each set  $X_t$  is an  $\mathcal{F}_t$ -measurable convex and closed polyhedron.

Suppose  $c_t$ ,  $t = 1, \dots, T$ , is an adapted sequence of random cost vectors.

A policy  $x$  results in a **cost sequence**  $Z_t = \langle c_t, x_t \rangle$ ,  $t = 1, \dots, T$ .

## Risk-averse multistage stochastic optimization problem

$$\min_{x \in I \cap F} \varrho(Z_1, Z_2, \dots, Z_T) \quad (\varrho - \text{dynamic measure of risk})$$

# Risk Evaluation on a Tree

## Scenario tree

Nodes  $n \in \mathcal{N}$ , organized in levels  $\Omega_t$  corresponding to stages  $1, \dots, T$ .

At level  $t = 1$  - only one **root node**  $n = 1$

Node  $n$  at level  $t$  is connected to an **ancestor node**  $a(n)$  at level  $t - 1$

Node  $n$  at level  $t$  is connected to a set  $C(n)$  of **children nodes** at  $t + 1$

## Value Function

$Q_n(x_{a(n)})$  - the best value of a subproblem rooted at node  $n$ , given  $x_{a(n)}$

$Q_C$  vector of value functions at nodes in the set  $C$

## Dynamic Programming Equations

$$Q_n(x_{a(n)}) = \min_{x_n} \left\{ \langle c_n, x_n \rangle : B_n x_{a(n)} + A_n x_n = b_n, x_n \in X_n \right\}, \quad n \in \Omega_T,$$

$$Q_n(x_{a(n)}) = \min_{x_n} \left\{ \langle c_n, x_n \rangle + \rho_n(Q_{C(n)}(x_n)) : \right. \\ \left. B_n x_{a(n)} + A_n x_n = b_n, x_n \in X_n \right\}, \quad n \in \Omega_t, \quad t = T - 1, \dots, 1$$

The optimal value functions  $Q_n(\cdot)$  are convex.

- State space  $\mathcal{X}$  (Borel)
- Control space  $\mathcal{U}$  (Borel)
- Feasible control set  $U : \mathcal{X} \rightrightarrows \mathcal{U}, t = 1, 2, \dots$
- Controlled transition kernel  $Q : \text{graph}(U) \rightarrow \mathcal{P}(\mathcal{X}), t = 1, 2, \dots$   
 $\mathcal{P}(\mathcal{X})$  - set of probability measures on  $\mathcal{X}$
- Cost functions  $c : \mathcal{X} \times \mathcal{U} \rightarrow \mathbb{R}, t = 1, 2, \dots$
- State history  $h_t = (x_1, \dots, x_t) \in \mathcal{X}^t$  (up to time  $t = 1, 2, \dots$ )
- Policy  $\pi_t : \mathcal{X}^t \rightarrow \mathcal{U}, t = 1, 2, \dots$  (always supported in  $U(x_t)$ )
- Markov policy  $\pi_t : \mathcal{X} \rightarrow \mathcal{U}, t = 1, 2, \dots$   
(stationary if  $\pi_t = \pi_1$  for all  $t$ )

$$x_t \longrightarrow u_t = \pi_t(x_t)$$

$$(x_t, u_t) \longrightarrow x_{t+1} \sim Q(x_t, u_t)$$

# Risk-Neutral Total Cost Problem

Infinite horizon expected cost problem:

$$\min_{\pi_1, \pi_2, \dots} \mathbb{E}^{\pi} \left[ \sum_{t=1}^{\infty} \alpha^{t-1} c_t(x_t, u_t) \right], \quad \alpha \in (0, 1]$$

with controls  $u_t = \pi_t(x_1, \dots, x_t)$

Two Cases:

Discounted models (with  $\alpha < 1$ ) and transient models (with  $\alpha = 1$ )

Standard Results:

- A **deterministic Markov policy** is optimal
- Optimal policy can be found by **dynamic programming equations**

Our Intention

Introduce **risk aversion** to the problem by replacing the expected value by **dynamic risk measures**

- Controlled Markov process  $x_t^\Pi$ ,  $t = 1, \dots, T$
- Policy  $\Pi = \{\pi_1, \pi_2, \dots, \pi_T\}$  with  $u_t = \pi_t(x_t)$  implies measure  $P^\Pi$
- Cost sequence  $Z_t^\Pi = c(x_t^\Pi, \pi_t(x_t^\Pi))$ ,  $Z_t \in \mathcal{Z}_t$ ,  $t = 1, \dots, T$ ,
- **Dynamic time-consistent risk measure**

$$J_T(\Pi) = Z_1^\Pi + \rho_1^\Pi(Z_2^\Pi + \dots + \rho_{T-1}^\Pi(Z_T^\Pi) \dots)$$

- Risk-averse optimal control problem

$$\min_{\Pi} \lim_{T \rightarrow \infty} J_T(\Pi)$$

## Difficulties

- Probability measure  $P^\Pi$ , processes  $x_t^\Pi$  and  $Z_t^\Pi$  depend on policy  $\Pi$
- The risk measures  $\rho_t^\Pi(\cdot)$  depend on  $\Pi$  and may depend on history; no Markov policies

- State space  $\mathcal{X}$  (Borel)
- Control space  $\mathcal{U}$  (Borel)
- State history  $h_t = (x_1, \dots, x_T) \in \mathcal{X}^t$  (up to time  $t = 1, 2, \dots$ )
- Controlled transition kernels  $Q_t : \mathcal{X}^t \times \mathcal{U} \rightarrow \mathcal{P}(\mathcal{X})$ ,  
 $\mathcal{P}(\mathcal{X})$  - set of probability measures on  $\mathcal{X}$
- Feasible control sets  $U_t : \mathcal{X}^t \rightrightarrows \mathcal{U}$ ,  $t = 1, 2, \dots$
- Cost functions  $c_t : \mathcal{X} \times \mathcal{U} \rightarrow \mathbb{R}$ ,  $t = 1, 2, \dots$
- Policy  $\pi_t : \mathcal{X}^t \rightarrow \mathcal{U}$ ,  $t = 1, 2, \dots$  (always supported in  $U_t(h_t)$ )

$$h_t \longrightarrow u_t = \pi_t(h_t)$$

$$(h_t, u_t) \longrightarrow x_{t+1} \sim Q_t(h_t, u_t) = Q_t^{\Pi}(h_t)$$

We only need to evaluate risk of processes  $Z_t^{\Pi}(h_t) = c(x_t, \pi_t(h_t))$ ,  
 $t = 1, \dots, T$ , which are **measurable functions of the history  $h_t$**

History  $h_t = (x_1, \dots, x_t)$ . Process  $Z_t^\Pi(h_t) = c(x_t, \pi_t(h_t))$ ,  $t = 1, \dots, T$

A family of conditional risk measures  $\{\rho_{t,T}^\Pi\}_{t=1,\dots,T}^{\pi \in \Pi}$  is **stochastically conditionally time-consistent** if for all feasible policies  $\Pi, \Pi'$ , all  $1 \leq t \leq T-1$ , and for all histories  $h_t, h'_t \in \mathcal{X}^t$ , the relations

$$Z_t^\Pi(h_t) = Z_t^{\Pi'}(h'_t)$$

$$(\rho_{t+1,T}^\Pi(Z_{t+1}^\Pi, \dots, Z_T^\Pi) | H_t^\Pi = h_t) \leq_{\text{st}} (\rho_{t+1,T}^{\Pi'}(Z_{t+1}^{\Pi'}, \dots, Z_T^{\Pi'}) | H_t^{\Pi'} = h'_t)$$

imply

$$\rho_{t,T}^\Pi(Z_t^\Pi, \dots, Z_T^\Pi)(h_t) \leq \rho_{t,T}^{\Pi'}(Z_t^{\Pi'}, \dots, Z_T^{\Pi'})(h'_t).$$

The conditional stochastic order  $\leq_{\text{st}}$ :

$$Q_t^\Pi(h_t)(\{y : Z_t^\Pi(h_t) + \rho_{t+1,T}^\Pi(Z_{t+1}^\Pi, \dots, Z_T^\Pi)(h_t, y) > \eta\})$$

$$\leq Q_t^{\Pi'}(h'_t)(\{y : Z_t^{\Pi'}(h'_t) + \rho_{t+1,T}^{\Pi'}(Z_{t+1}^{\Pi'}, \dots, Z_T^{\Pi'})(h'_t, y) > \eta\})$$

The processes evaluated are  $Z_t^\Pi(h_t) = c(x_t, \pi_t(h_t))$ ,  $t = 1, \dots, T$

A family of dynamic risk measures  $\{(\rho_{t,T}^\Pi)_{t=1,\dots,T} : \Pi \in \Pi\}$  is translation-invariant and stochastically conditionally time-consistent if and only if there exist functionals

$$\sigma_t : \mathcal{V} \times (\cup_{\Pi \in \Pi} \text{Graph}(Q_t^\Pi)) \rightarrow \mathbb{R}, \quad t = 1 \dots T - 1,$$

where  $\mathcal{V}$  is the set of measurable functions on  $(\mathcal{X}, \mathcal{B}(\mathcal{X}))$ , and

$$\rho_{t,T}^\Pi(Z_t^\Pi, \dots, Z_T^\Pi)(h_t) = Z_t^\Pi(h_t) + \sigma_t(\rho_{t+1,T}^\Pi(Z_{t+1}^\Pi, \dots, Z_T^\Pi)(h_t, \cdot), h_t, Q_t^\Pi(h_t))$$

For all  $\Pi \in \Pi$ ,  $h_t \in \mathcal{X}^t$ , the function  $\sigma_t(\cdot, h_t, Q_t^\Pi(h_t))$  is a **law-invariant risk measure** on  $(\mathcal{X}, \mathcal{B}(\mathcal{X}), Q_t^\Pi(h_t))$ .

The mapping  $\sigma_t$  does not depend on  $\Pi$ : the policy only affects the equation through the next state's distribution  $Q_t^\Pi(h_t)$ .



A family of process-based dynamic risk measures  $\{\rho_{t,T}^{\Pi}\}_{t=1,\dots,T}^{\Pi \in \Pi}$  for a Markov decision problem is **Markov** if for all Markov policies  $\Pi \in \Pi$ , for any measurable  $c_1, \dots, c_T : \mathcal{X} \times \mathcal{U} \rightarrow \mathbb{R}$ , and for all  $h_t = (x_1, \dots, x_t)$  and  $h'_t = (x'_1, \dots, x'_t)$  such that  $x_t = x'_t$ , we have

$$\begin{aligned} \rho_{t,T}^{\Pi}(c_t(X_t, \pi_t(X_t)), \dots, c_T(X_T, \pi_T(X_T)))(h_t) \\ = \rho_{t,T}^{\Pi}(c_t(X_t, \pi_t(X_t)), \dots, c_T(X_T, \pi_T(X_T)))(h'_t). \end{aligned}$$

If the current state  $x_t$  is the same, and the same Markov policy  $\Pi$  is used, then the risk is the same.

For a fixed history-dependent policy  $\Pi$  and every  $h_t \in \mathcal{X}^t$ , we write

$$v_t^{c, \Pi}(h_t) = \rho_{t, T}^{\Pi}(c_t(X_t, \pi_t(H_t)), \dots, c_T(X_T, \pi_T(H_T)))(h_t)$$

If a family of process-based dynamic risk measures  $\{\rho_{t, T}^{\Pi}\}_{t=1, \dots, T}^{\Pi \in \Pi}$  is Markov, translation-invariant, and stochastically conditionally time-consistent, then there exist **transition risk mappings**

$$\sigma_t : \mathcal{V} \times \{(x, Q_t(x, u)) : u \in U(x), x \in \mathcal{X}\} \rightarrow \mathbb{R}, \quad t = 1, \dots, T - 1$$

such that for all  $\Pi \in \Pi$ , for all  $t = 1, \dots, T - 1$ , and all  $h_t \in \mathcal{X}^t$ , the functional  $\sigma_t(\cdot, x_t, Q_t(x_t, \pi_t(h_t)))$  is a **law-invariant risk measure** on  $(\mathcal{X}, \mathcal{B}(\mathcal{X}), Q_t(x_t, \pi_t(h_t)))$ . Moreover, for any  $c = \{c_t\}_{t=1 \dots T}$ , we have

$$v_t^{c, \Pi}(h_t) = c_t(x_t, \pi_t(h_t)) + \sigma_t(v_{t+1}^{c, \Pi}(h_t, \cdot), x_t, Q_t(x_t, \pi_t(h_t))), \quad t = 1, \dots, T - 1$$

From now on we assume that  $\sigma_t(\cdot, x, m)$  is a **coherent risk measure** on  $\mathcal{V} = \mathcal{L}_p(\mathcal{X}, \mathcal{B}, P_0)$ .

## Dual representation of transition risk mappings

$$\sigma(v, x, m) = \max_{\mu \in \mathcal{A}(x, m)} \int_{\mathcal{X}} v(y) \mu(dy)$$

## Example: Mean–Semideviation

$$\sigma(v, x, m) = \int v dm + \kappa(x) \left( \int (v - \int v dm)_+^p dm \right)^{\frac{1}{p}}$$

For  $p > 1$  we obtain

$$\mathcal{A}(x, m) = \left\{ g = m \left( 1 + h - \int h dm \right) : \|h\|_{\mathcal{L}_q(\mathcal{X}, \mathcal{B}, m)} \leq \kappa(x), h \geq 0 \right\}$$

- G0.** For all  $x \in \mathcal{X}$ ,  $u \in U_t(x)$  the measure  $Q_t(x, u)$  is an element of  $\mathcal{V}'$ ;
- G1.** The transition kernel  $Q_t(\cdot, \cdot)$  is setwise continuous;
- G2.** The multifunctions  $\mathcal{A}_t(\cdot, \cdot) \equiv \partial_\varphi \sigma_t(0, \cdot, \cdot)$  are lower semicontinuous;
- G3.** The functions  $c_t(\cdot, \cdot)$  are measurable,  $w$ -bounded, and lower semicontinuous;
- G4.** The multifunctions  $U_t(\cdot)$  are measurable and compact-valued.

## Finite Horizon Risk-Averse Control Problem

Consider a controlled Markov process  $\{x_t\}$  with  $u_t = \pi_t(x_1, \dots, x_t)$ .

Risk-averse optimal control problem:

$$\min_{\Pi} J_T(\Pi, x_1) = c_1(x_1, u_1) + \rho_1^{\Pi} \left( c_2(x_2, u_2) + \dots \right. \\ \left. + \rho_{T-1}^{\Pi} \left( c_T(x_T, u_T) + \rho_T(c_{T+1}(x_{T+1})) \dots \right) \right).$$

### Theorem

If the conditional measures  $\rho_t^{\Pi}$  are Markov (+ general conditions), then the optimal solution is given by the **dynamic programming equations**:

$$v_{T+1}(x) = c_{T+1}(x), \quad x \in \mathcal{X}$$

$$v_t(x) = \min_{u \in U(x)} \left\{ c_t(x, u) + \sigma_t(v_{t+1}, x, Q_t(x, u)) \right\}, \quad x \in \mathcal{X}, \quad t = T, \dots, 1.$$

Optimal **Markov policy**  $\hat{\Pi} = \{\hat{\pi}_1, \dots, \hat{\pi}_T\}$  - the minimizers above

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$$v_{T+1}(x) = c_{T+1}(x), \quad x \in \mathcal{X}$$

$$v_t(x) = \min_{u \in U(x)} \left\{ c_t(x, u) + \max_{\mu \in \mathcal{A}_t(x, Q_t(x, u))} \mathbb{E}_{\mu} [v_{t+1}] \right\}, \quad x \in \mathcal{X}, \quad t = T, \dots, 1.$$

Optimal **Markov policy**  $\hat{\Pi} = \{\hat{\pi}_1, \dots, \hat{\pi}_T\}$  - the minimizers above

## Infinite Horizon Risk (for stationary models)

Discounted risk measure ( $0 < \alpha < 1$ )

$$J_T^\alpha(\Pi, x) = Z_1^\Pi + \rho_1^\Pi \left( \alpha Z_2^\Pi + \dots + \rho_{T-1}^\Pi \left( \alpha^{T-1} Z_T^\Pi \right) \dots \right)$$

Optimal cost:  $J^*(x) = \inf_{\Pi} \lim_{T \rightarrow \infty} J_T^\alpha(\Pi, x)$

Assume that the model is stationary, the conditional risk measures  $\rho_t$ ,  $t = 1, \dots, T$ , are **Markov** (+ technical conditions). Then a bounded function  $v : \mathcal{X} \rightarrow \mathbb{R}$  satisfies the **dynamic programming equations**

$$v(x) = \min_{u \in U(x)} \left\{ c(x, u) + \alpha \sigma(v, x, Q(x, u)) \right\}, \quad x \in \mathcal{X},$$

if and only if  $v(\cdot) \equiv J^*(\cdot)$ . Moreover, the minimizer  $\pi^*(x)$ ,  $x \in \mathcal{X}$ , on the right hand side exists and defines an **optimal Markov policy**  $\Pi^* = \{\pi^*, \pi^*, \dots\}$ .

If  $\alpha = 1$  additional conditions of **risk transient models**

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$$v(x) = \min_{u \in U(x)} \left\{ c(x, u) + \alpha \max_{\mu \in \mathcal{A}(x, Q(x, u))} \mathbb{E}_\mu[v] \right\}, \quad x \in \mathcal{X},$$

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If  $\alpha = 1$  additional conditions of **risk transient models**



For every  $x$  we define the set of probability measures:

$$\mathfrak{M}^\pi(x) = \mathcal{A}(x, Q(x, \pi(x))), \quad x \in \mathcal{X}$$

The multifunction  $\mathfrak{M}^\pi : \mathcal{X} \rightrightarrows \mathcal{P}(\mathcal{X})$  is a **risk multikernel**, associated with the risk transition mapping  $\sigma(\cdot, \cdot, \cdot)$ , the kernel  $Q$ , and decision rule  $\pi$ .

Key formula for Markov policy  $\Pi = \{\pi, \pi, \dots\}$

$$\rho_t^\Pi(v(x_{t+1})) = \max_{M \in \mathfrak{M}^\pi(x_t)} \int_{\mathcal{X}} v(y) M(dy)$$

A Markov model is **risk-transient** if

$$\|M\|_w \leq K \quad \text{for all } M \ll \sum_{j=1}^T (\tilde{\mathfrak{M}}^\pi)^j \quad \text{and all } T \geq 0$$

## Value iteration

$$v^{k+1}(x) = \min_{u \in U(x)} \left\{ c(x, u) + \alpha \sigma(v^k, x, Q(x, u)) \right\}, \quad x \in \mathcal{X}, \quad k = 1, 2, \dots$$

## Policy iteration

- For  $k = 0, 1, 2, \dots$ , given a stationary Markov policy  $\{\pi^k, \pi^k, \dots\}$ , find the **value function**  $v^k$  by solving (by a specialized Newton method) the **nonsmooth equation**

$$v(x) = c(x, \pi^k(x)) + \alpha \sigma(v, x, Q(x, \pi^k(x))), \quad x \in \mathcal{X}$$

- Find the **next policy**  $\pi^{k+1}(\cdot)$  by **one-step optimization**

$$\pi^{k+1}(x) = \operatorname{argmin}_{u \in U(x)} \left\{ c(x, u) + \alpha \sigma(v^k, x, Q(x, u)) \right\}, \quad x \in \mathcal{X}$$

For  $\alpha = 1$  additional conditions of **risk transient models**  
+ positive or negative  $c(\cdot, \cdot)$  for the value iteration method

Offers  $Y_t$  arriving in time periods  $t = 1, 2, \dots$  are i.i.d. integrable random variables. At each time we may accept the highest offer so far, or wait, at cost  $c_0$ .

**The expected value solution:** accept the first offer greater than or equal to the solution  $\hat{x}$  of the equation

$$\mathbb{E}[(Y - \hat{x})_+] = c_0.$$

**Risk-averse DP equation:**

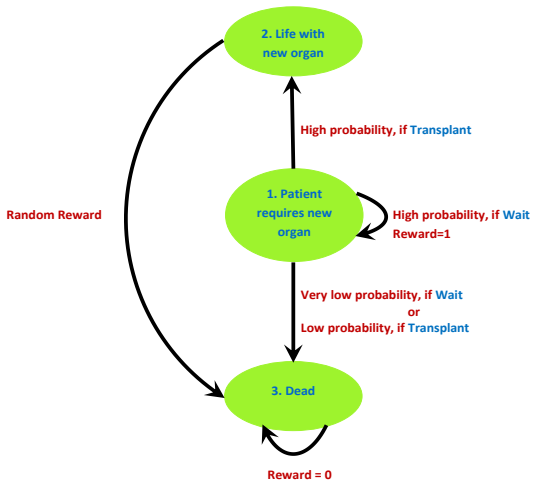
$$v(x) = \min \left\{ -x, c_0 + \sigma(v, x, Q(x)) \right\}, \quad x \in \mathbb{R}_+$$

Suppose  $\sigma$  is law invariant and does not depend on the second argument.

**Risk-averse solution:** accept any offer that is greater or equal to the solution  $x^*$  of the equation

$$c_0 = \min_{\mu \in \mathcal{A}} \mathbb{E}_{\mu}[(Y - x^*)_+] \quad (\mathcal{A} - \text{subdifferential of } \sigma).$$

If  $x < x^*$ , then wait.



- **Expected Total Reward:**

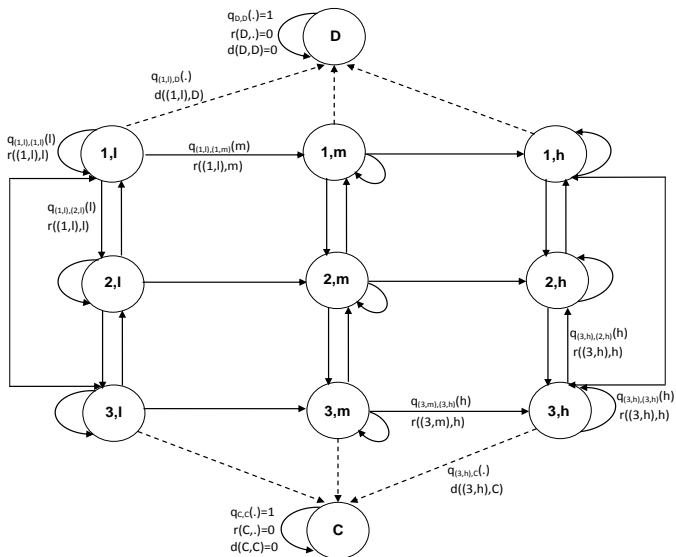
The optimal policy is to wait

- **Mean Semi-Deviation with Deterministic Policies:**

The optimal policy is to transplant

- **Mean Semi-Deviation with Randomized Policies:**

Wait with probability 0.993983 and transplant with probability 0.006017



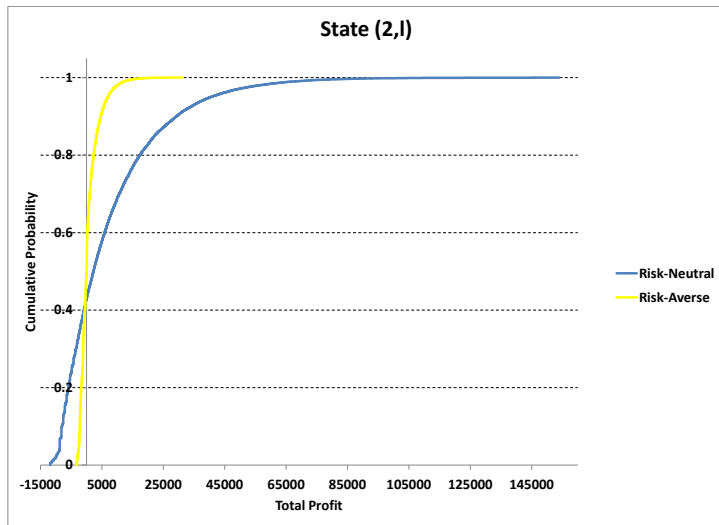
$\kappa$	(1,l)	(1,m)	(1,h)	(2,l)	(2,m)	(2,h)	(3,l)	(3,m)	(3,h)
0.025	m	h	h	m	h	h	m	m	h
0.1	l	h	h	m	h	h	m	m	h
0.2	l	h	h	m	h	h	m	m	h
0.3	l	h	h	m	h	h	m	m	h
0.4	l	h	h	m	h	h	m	m	h
0.5	l	h	h	m	h	h	m	m	h
0.6	l	h	h	m	h	h	m	m	h
0.7	l	h	h	m	h	h	m	m	h
0.8	l	m	h	l	h	h	m	m	h
0.9	l	m	h	l	m	h	m	m	h
1	l	m	h	l	m	h	m	m	h

$\kappa$  - risk aversion coefficient

# Comparison of Methods

$\kappa$	# of Value Iterations	# of Policy Iterations	# of Newton Iterations
0.025	869	3	4,3,3
0.1	797	4	3,3,2,3
0.2	746	4	3,3,2,2
0.3	689	4	4,2,2,2
0.4	658	4	4,2,2,2
0.5	661	4	4,2,2,2
0.6	761	3	4,3,3
0.7	893	3	4,2,3
0.8	525	3	4,3,2
0.9	1354	3	5,2,3
1	1231	3	6,2,3





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